

ScPoEconometrics

Intro To Causality

Mylène Feuillade, Gustave Kenedi, Florian Oswald and Pierre Villedieu SciencesPo Paris 2022-02-22

Quick "Quiz" on Last Week's Material

1. From your *computer* (3) connect to *www.wooclap.com/SCPOSLR*OR

2. From your *phone* (F) flash QR code below





Today - Introduction to Causal Inference

- Causality versus correlation
- The *Potential Outcomes Framework* a.k.a. Rubin's Causal Model
- Randomized controlled trials (RCTs)
- Follow up on the empirical application of *class size* and *student performance*



Causality and Economics

- Making causal inference from data can be seen as economists' *comparative advantage* among the social sciences!
- Plenty of fields do statistics. But very few make it standard training for their students to understand causality.
- Economists' endeavour to establish causal relationships is also what makes them useful both in the private (e.g. tech companies) and public sector (e.g. policy advisors).



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- Economists' endeavour to establish causal relationships is also what makes them useful both in the private (e.g. tech companies) and public sector (e.g. policy advisors).
- Ok, that's enough preaching



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- \triangle It does **NOT** mean that X is the only factor that causes Y.



Correlation vs Causation

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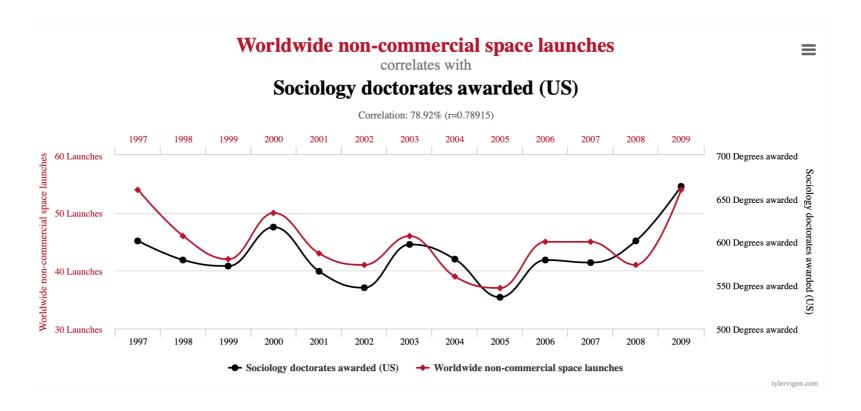
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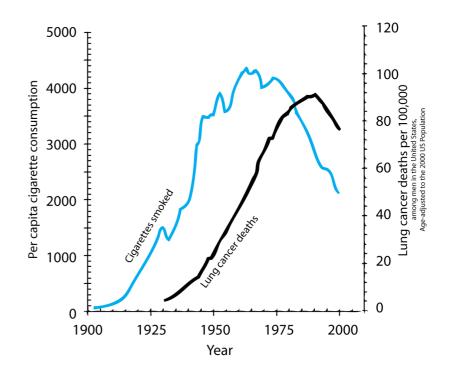
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 - We are at the start of a big increase in deaths from lung cancer...
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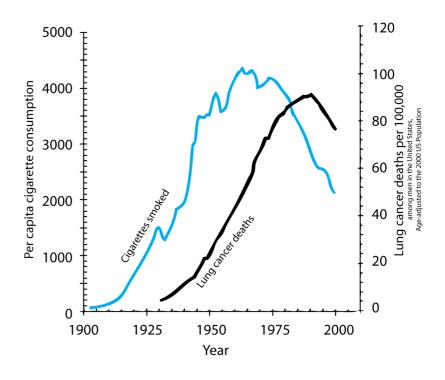




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• It's very tempting to claim that smoking causes lung cancer based on this graph.

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Macro confouding factors:

Other macro factors which can cause cancers also changed between 1900 and 1950:

- Tarring of roads,
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Self selection:

Smokers and non-smokers may be different in the first place:

- Selection on observable characteristics: age, education, income, etc.
- Selection on unobservable characteristics: genes (the hypothetical confounding genome theory of Fisher).



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Why might the observed correlation between *economic growth* and *financial development* not reflect the causal effect of the financial sector?

Again, maybe economic growth leads to financial development and not the other way around \rightarrow reverse causality / simultaneity



Link with Economic Theory

- Economic theory tells us individuals' behave in order to *maximise their utility*
- Thus they don't choose to act in $random\ ways \rightarrow$ we say that individual's behavior is endogenous
- We should be *suspicious* of any correlation found in data



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- Thus they don't choose to act in $random\ ways \rightarrow$ we say that individual's behavior is endogenous
- We should be *suspicious* of any correlation found in data
- How can we make *causal claims* then?
- The *Potential Outcomes Framework* will be our guide.



Causal Inference

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For practicality, let this treatment variable D_i be a binary variable:

$$D_i = \left\{ egin{aligned} 1 & ext{if individual } i ext{ is treated} \ 0 & ext{if individual } i ext{ is not treated} \end{aligned}
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Treatment group

all the individuals such that $D_i = 1$.

Control group

all the individuals such that $D_i=0$.



- In this framework, each individual has two **potential outcomes**, but only one **observed** outcome Y_i :
 - $\circ Y_i^1$: potential outcome if individual i receives the treatment $(D_i=1)$,
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- In real life we only observe Y_i which can be written as:

$$Y_i = D_i imes Y_i^1 + (1-D_i) imes Y_i^0$$



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• *Fundamental Problem of Causal Inference*: for any individual *i*, we only observe one of either potential outcomes (Holland, 1986).



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• From these we can define the *individual treatment effect* δ_i :

$$\delta_i = Y_i^{\,1} - Y_i^{\,0}$$

• δ_i measures the **causal effect of the treatment** (D_i) on outcome Y for individual i.



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- δ_i measures the **causal effect of the treatment** (D_i) on outcome Y for individual i.
- Since the treatment effect *cannot* be observed at the individual level, we estimate population averages.



Sidenote: Expectation and Conditional Expectation

Let's say you have a fair die and you roll it an infinite number of times. What is the average number rolled?

• If *X* is a random variable containing the number rolled, we write:

$$\mathbb{E}(X) = rac{1}{6} imes 1 + rac{1}{6} imes 2 + rac{1}{6} imes 3 + rac{1}{6} imes 4 + rac{1}{6} imes 5 + rac{1}{6} imes 6 = 3.5$$

- The $\mathbb{E}(.)$ operator stands for **expectation** or *population mean*.
- The $\mathbb{E}(.)$ operator is linear, in other words, $\mathbb{E}(X+Y)=\mathbb{E}(X)+\mathbb{E}(Y)$ with X and Y being two random variables.



Sidenote: Expectation and Conditional Expectation

Now, let's say you have two fair dice and you roll them an infinite number of times. What is the average sum of numbers rolled, conditional on one of them being always equal to 5?

• If X is a random variable containing the number rolled of die 1 and Y a random variable containing the number rolled of die 2, we write:

$$\mathbb{E}(X+Y|Y=5) = \mathbb{E}(X|Y=5) + \mathbb{E}(Y|Y=5)$$
 $= \mathbb{E}(X) + 5$
 $= 3.5 + 5$
 $= 8.5$

• The $\mathbb{E}(.\,|D=x)$ operator stands for **conditional expectation**. It refers to the expectation over a subcategory of the entire population, namely people who satisfy the condition D=x.



Average Treatment Effect (ATE)

Broadest possible average effect:

$$egin{aligned} ATE &= \mathbb{E}(\delta_i) \ &= \mathbb{E}(Y_i^1 - Y_i^0) \ &= \mathbb{E}(Y_i^1) - \mathbb{E}(Y_i^0) \end{aligned}$$

• The ATE simply measures the *average of individual treatment effects over the whole population*.

(Appendix: Average Treatment on the Treated and Average Treatment on the Untreated)



Example: Small vs. Large Class

Potential outcomes for students of being in a small (Y^1) or large class (Y^0) on GPA (0-10):

Student	Y^1	Y^0	δ
1	5	2	3
2	6	4	2
3	3	6	-3
4	5	4	1
5	10	8	2
6	2	4	-2
7	5	2	3
8	6	4	2
9	2	9	-7
10	8	2	6
Average	5.2	4.5	0.7



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$$egin{aligned} \mathbf{ATE} &= \mathbb{E}(\delta) \ &= \mathbb{E}(Y^1) - \mathbb{E}(Y^0) \ &= 5.2 - 4.5 \ &= 0.7 \end{aligned}$$

 \rightarrow the *average* causal effect of being in small relative to large class on GPA is 0.7 points.

♠ not all students benefited equally from the treatment!



The Problem of Causal Inference

• In practice, we have the same **missing data problem** for computing the ATE as we did for δ_i . Either Y_i^0 or Y_i^0 is missing for each i.



The Problem of Causal Inference

- In practice, we have the same **missing data problem** for computing the ATE as we did for δ_i . Either Y_i^1 or Y_i^0 is missing for each i.
- From the data, we can compute the Simple Difference in mean Outcomes (SDO) for both groups:

$$SDO = \mathbb{E}(Y_i^1|D_i=1) - \mathbb{E}(Y_i^0|D_i=0) \ = \underbrace{\frac{1}{N_T}\sum_{i=1}^{N_T}(Y_i|D_i=1)}_{ ext{average outcome of treatment group}} - \underbrace{\frac{1}{N_C}\sum_{i=1}^{N_C}(Y_i|D_i=0)}_{ ext{average outcome of control group}}$$



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Average			0.7

The simple difference in mean outcomes:

$$SDO = rac{5+6+5+10+5+6+8}{7} - rac{6+4+9}{3} \ pprox 6.43-6.33 pprox 0.1$$

- The SDO is much smaller than the ATE!
- Such a difference will (almost always)
 fail to capture the causal treatment
 effect
- Notice that this kind "naive" comparison is often done by journalists, politicians, badly trained scientists (but not you now!



Problems with Naive Comparisons

Let's rewrite the SDO to make the individual treatment effect (δ_i) appear in the equation.

$$egin{aligned} SDO &= \mathbb{E}(Y_i^1|D_i = 1) - \mathbb{E}(Y_i^0|D_i = 0) \ &= \mathbb{E}(Y_i^0 + \delta_i|D_i = 1) - \mathbb{E}(Y_i^0|D_i = 0) \end{aligned}$$



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Then,

$$SDO = \delta + \mathbb{E}(Y_i^0|D_i=1) - \mathbb{E}(Y_i^0|D_i=0)$$

And because $ATE = \mathbb{E}(\delta_i) = \mathbb{E}(\delta) = \delta$ (by assumption), we get:

$$SDO = ATE + \underbrace{\mathbb{E}(Y_i^0|D_i=1) - \mathbb{E}(Y_i^0|D_i=0)}_{ ext{Selection bias}}$$



(Appendix: when constant treatment assumption is relaxed another bias term appears.)

Task 1: SDO, ATE and Randomization

Let's compute these various quantities and biases with data we generated ourselves.

- 1. Load the data here using read.csv. The group variable corresponds to whether the individual has been treated or not, Y0 to the potential outcome if the individual does not receive the treatment (Y_i^0) while Y1 to the potential outcome if the individual receives the treatment (Y_i^1) . Create a new variable containing the observed outcome (Y_i) and the individual treatment effect (δ_i) . Recall $Y_i = D_i \times Y_i^1 + (1 D_i) \times Y_i^0$, $\delta_i = Y_i^1 Y_i^0$.
- 2. Compute the *ATE* and the *SDO*. Is there is any *bias*? Is it large in magnitude?
- 3. In this new dataset we've randomly assigned the same individuals to the treatment and control groups. Compute the *SDO under randomization*. Remember that you need to recompute Y_i because the assignment is not the same anymore. If you got it right, the bias should be very close to 0. Why is it not exactly 0?
- 4. Optional: Compute the selection bias and the heterogeneous treatment effect bias and check that SDO = ATE + selection bias + heterogeneous treatment effect bias



Randomization solves the problem of causal inference!

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 - Therefore the *selection bias is equal to 0*.
- With random assignment we have:

$$SDO = \mathbb{E}(Y_i^1|D_i=1) - \mathbb{E}(Y_i^0|D_i=0) = ATE$$

⟨ We can directly estimate the ATE from the data!



Randomized Experiments

Randomized Experiments

- Often called Randomized Controlled Trials (RCT).
- The first RCTs were conducted a long time ago (18th and 19th century), mainly in **Medecine**.
- In the beginning of the 20th century they were popularized by famous statisticians like **J. Neyman** or **R.A. Fisher**.
- Since then they have had a growing influence and have progressively become a reliable tool for public policy evaluation.
- As for economics, the **2019 Nobel Price in Economics** was awarded to three exponents of RCTs, Abhijit Banerjee, Esther Duflo and Michael Kremer, "for their experimental approach to alleviating global poverty".



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$$\operatorname{math} \operatorname{score}_i = b_0 + b_1 \operatorname{class} \operatorname{size}_i + e_i$$



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We briefly discussed why b_1^{OLS} could only establish an **association** and not a **causal** relationship.

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• There was a problem of *non-random attrition* but we will ignore it.

Task 2: STAR data

- 1. Load the *STAR* data from here and assign it to an object called <code>star_df</code>. Read the help for the data here to understand what the variables correspond to. (Note: the data has been *reshaped* so don't mind the "k", "1", etc. in the variable names in the help.)
- 2. What's the unit of observation? Which variable contains: (i) the (random) class assignment, (ii) the student's class grade, (iii) the outcomes of interest?
- 3. How many observations are there? Why so many if 11,598 students participated? Why are there so many NAs? What do they correspond to?
- 4. Keep only cases with no NAs with the following code: star_df <- star_df[complete.cases(star_df),]
- 5. Let's check how well the randomization was done by doing *balancing checks*. Compute the average percentage of girls, african americans, and free lunch qualifiers by grade *and* treatment class. (*Hint*: The following computes the percentage of girls (without the relevant dplyr verbs): share_female = mean(gender == "female") * 100.)



We just saw that in an RCT the Average Treatment Effect is obtained by computing the differences in outcomes between the treatment and control groups.

Let's only focus on:

- One treatment group: small classes,
- One control group: regular classes,
- One grade: **kindergarten** (k).



The Project STAR Experiment

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grade	test	mean regular	mean small	ATE
k	math	484.45	493.34	8.9
k	read	435.76	441.13	5.37

What's the interpretation for these ATEs?



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What's the interpretation for these ATEs?



That's nice but can't we put this in regression form?

$$Y_i = D_i Y_i^{\, 1} + (1 - D_i) Y_i^{\, 0}$$



$$Y_i = D_i Y_i^1 + (1 - D_i) Y_i^0$$

Factoring by D_i and replacing $Y_i^{\, 1} - Y_i^{\, 0}$ by δ_i , we get:

$$egin{aligned} Y_i &= Y_i^0 + D_i (Y_i^1 - Y_i^0) \ &= Y_i^0 + D_i \delta_i \end{aligned}$$



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Assuming $\delta_i = \delta$, for all i,

$$Y_i = Y_i^{\,0} + D_i \delta$$



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Adding $\mathbb{E}[Y_i^0] - \mathbb{E}[Y_i^0] = 0$ to the right-hand side:

$$egin{aligned} Y_i &= \mathbb{E}[Y_i^0] + D_i \delta + Y_i^0 - \mathbb{E}[Y_i^0] \ &= b_0 + \delta D_i + e_i \end{aligned}$$



The last equation looks exactly like the simple regression model we saw last week! (with $\delta=b_1$)

Let's therefore estimate the ATE of being assigned to a small class size on math scores.



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Let's therefore estimate the ATE of being assigned to a small class size on math scores.

We want to estimate the following model: math $\mathrm{score}_i = b_0 + \delta \mathrm{small}_i + e_i$, with

$$\text{small}_i = \left\{ egin{aligned} 1 & \text{if assigned to a small class} \\ 0 & \text{if assigned to a regular class} \end{array}
ight.$$

```
## # A tibble: 2 x 3
star_df_k_small <- star_df %>%
                                                     ## star
                                                                 grade
   filter(star %in% c("regular", "small") &
                                                     ## <chr> <chr> <int>
          grade == "k") %>%
                                                     ## 1 regular k
                                                                         1781
 mutate(small = (star == "small"))
                                                     ## 2 small
                                                                        1578
                                                     ## # A tibble: 2 x 2
star_df_k_small %>% count(star, grade)
                                                          small
                                                          <lgl> <int>
star df k small %>% count(small)
                                                     ## 1 FALSE
                                                               1781
                                                     ## 2 TRUE
                                                                 1578
```



Regression model we want to estimate: math $\mathrm{score}_i = b_0 + \delta \mathrm{small}_i + e_i$

```
lm(math ~ small, star_df_k_small)

##
## Call:
## lm(formula = math ~ small, data = star_df_k_small)
##
## Coefficients:
## (Intercept) smallTRUE
## 484.446 8.895
```



Regression model we want to estimate: $\mathrm{math}\ \mathrm{score}_i = b_0 + \delta \mathrm{small}_i + e_i$

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##
## Call:
## lm(formula = math ~ small, data = star_df_k_small)
##
## Coefficients:
## (Intercept) smallTRUE
## 484.446 8.895
```

```
Recall that: b_0 = \mathbb{E}[Y_i^0] and \delta = \mathbb{E}[Y_i|D_i=1] - \mathbb{E}[Y_i|D_i=0]
```

```
b_0 = mean(star_df_k_small$math[
    star_df_k_small$small == FALSE])
b_0

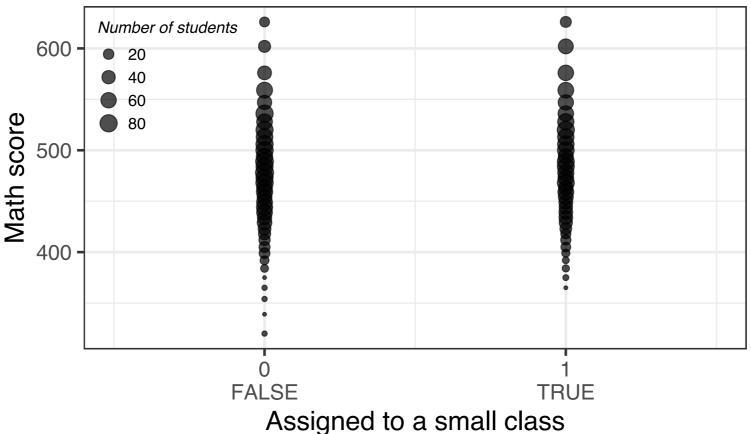
## [1] 484.4464

delta = mean(star_df_k_small$math[
    star_df_k_small$small == TRUE]) -
    mean(star_df_k_small$math[
    star_df_k_small$small == FALSE])
    delta

## [1] 8.895193
```

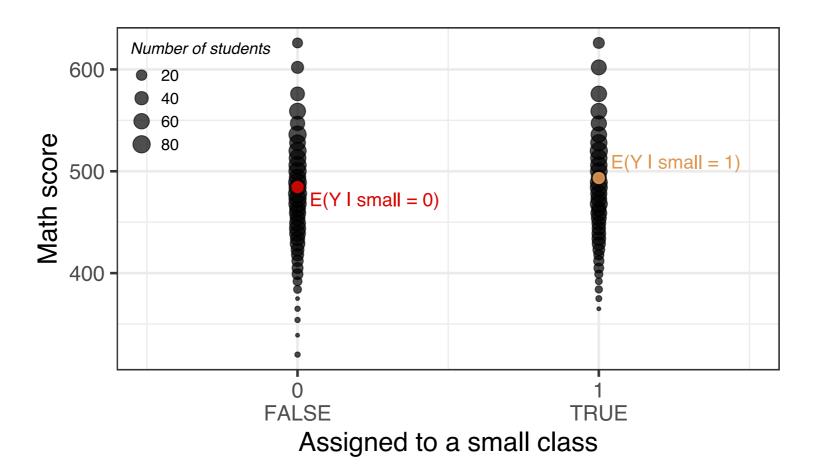


Contrary to last week, the regressor in our regression is a *dummy variable*, i.e. a variable that takes the values TRUE or FALSE (1 or 0).



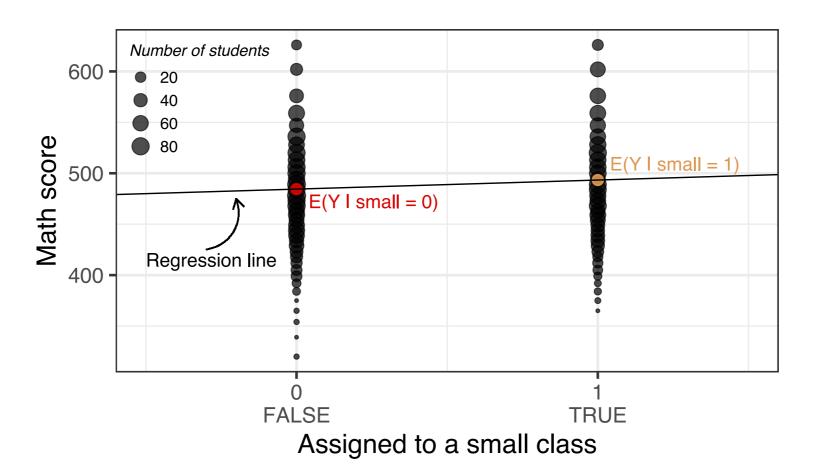


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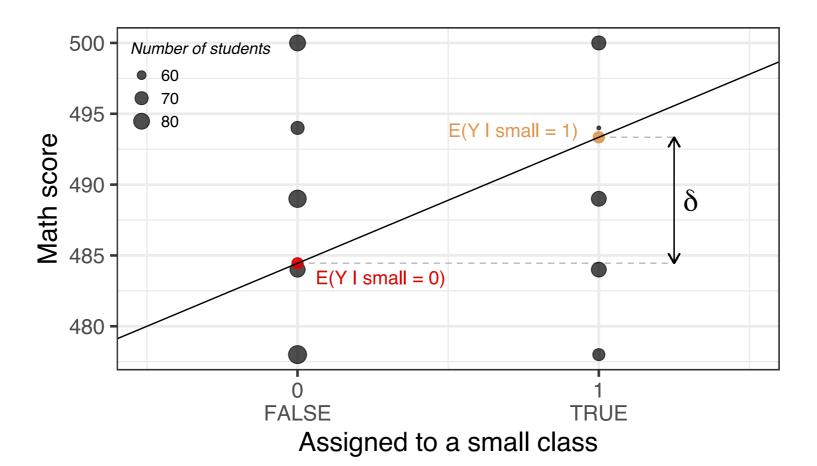


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Recall the regression model: math $\mathrm{score}_i = b_0 + \delta \mathrm{small}_i + e_i$

$$egin{aligned} \mathbb{E}[ext{math score}| ext{small}_i = 0] &= \mathbb{E}[b_0 + \delta ext{small}_i + e_i| ext{small}_i = 0] \ &= b_0 + \delta \mathbb{E}[ext{small}_i| ext{small}_i = 0] + \mathbb{E}[e_i| ext{small}_i = 0] \ &= b_0 \end{aligned}$$



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$$egin{aligned} ATE &= \mathbb{E}[ext{math score}| ext{small}_i = 1] - \mathbb{E}[ext{math score}| ext{small}_i = 0] \ &= b_0 + \delta - b_0 \ &= \delta \end{aligned}$$



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We knew this already but we now understand why this is true 🖞

Task 3: Your Turn!

Run the following code to filter the dataset to only keep first graders and the small class and regular class groups:

```
star_df_clean <- star_df %>%
  filter(grade == "1" & star %in% c("small", "regular"))
```

- 1. Compute the average math score for both groups, and the difference between the two. (Use base R.)
- 2. Create a dummy variable treatment equal to TRUE if student is in treatment group (i.e. small class size) and FALSE if in control group (i.e. regular class size). *Hint:* you can create the dummy variable with treatment = (star == "small").
- 3. Regress math score on the treatment dummy variable. Are the results in line with question 2?
- 4. How do you interpret these coefficients?



Shortcomings of RCTs

RCTs have very strong *internal validity*, that is they can convincingly establish causal links.

However, they have some shortcomings:

- RCT are often **infeasible**:
 - RCTs are **costly**,
 - RCTs may face some **ethical issues**: some *treatments* simply cannot be given to people,
 - RCTs take time and we may be **time constrained**.



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However, they have some shortcomings:

- RCT are often **infeasible**:
 - RCTs are costly,
 - RCTs may face some ethical issues: some treatments simply cannot be given to people,
 - RCTs take time and we may be time constrained.
- **Interpretation** of the results:
 - *External validity*: To what extent can the results from a given RCT be generalized to other contexts (countries, populations,...)?
 - Uncovering the mechanisms that are at stake may be difficult,
 - Imperfect randomization, attrition, ...



What comes next?

• So if we cannot rely on RCTs to make our life easy, it means we have to find a way to make causal inference from *observational data* (as opposed to *experimental data*).



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• So if we cannot rely on RCTs to make our life easy, it means we have to find a way to make causal inference from *observational data* (as opposed to *experimental data*).

2 broad cases:

- selection occurs on observable characteristics: multiple regression (next week!)
- selection occurs on unobservable characteristics: regression discontinuity design (lecture 10) or difference-in-differences (not covered but set of slides online!)



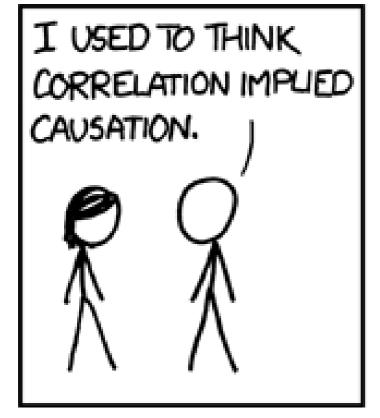
On the way to causality

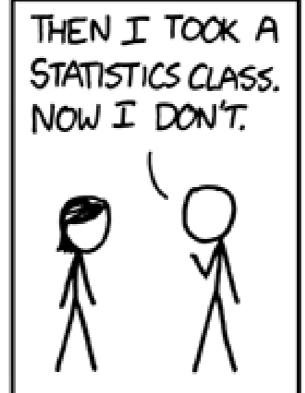
☑ How to manage data? Read it, tidy it, visualise it...

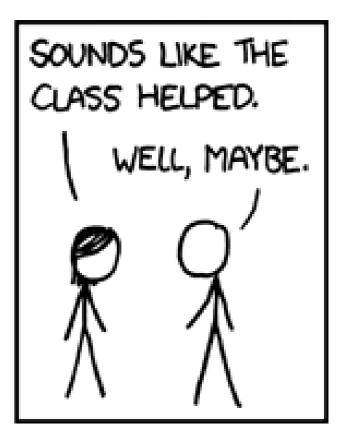
How to summarise relationships between variables? Simple linear regression... to be continued

- **What is causality?**
- X What if we don't observe an entire population?
- X Are our findings just due to randomness?
- **X** How to find exogeneity in practice?













SEE YOU NEXT WEEK!







@ScPoEcon



Appendix

Average Treatment on the Treated and on the Untreated

Other *conditional* average treatment effects may be of interest:

Average Treatment on the Treated (ATT)

$$egin{aligned} ATT &= \mathbb{E}(\delta_i | D_i = 1) \ &= \mathbb{E}(Y_i^{\ 1} - Y_i^{\ 0} | D_i = 1) \ &= \mathbb{E}(Y_i^{\ 1} | D_i = 1) - \mathbb{E}(Y_i^{\ 0} | D_i = 1) \end{aligned}$$

The ATT measures the average treatment effect conditional on being in the treatment group.

Example: the effect of participating in a training program (treatment) for those who participated (treatment group).

Average Treatment on the Untreated (ATU)

$$egin{aligned} ATU &= \mathbb{E}(\delta_i | D_i = 0) \ &= \mathbb{E}(Y_i^1 - Y_i^0 | D_i = 0) \ &= \mathbb{E}(Y_i^1 | D_i = 0) - \mathbb{E}(Y_i^0 | D_i = 0) \end{aligned}$$

The ATU measures the average treatment effect conditional on being in the control group.

Example: the effect of attending a private school (*treatment*) for students from a public school (*control group*).



Problems with Naive Comparisons

Let's now relax the assumption that the treatment effect is constant among all individuals.

After some tedious calculations that we skip, the SDO can now be decomposed as:

$$SDO = ATE + \underbrace{\mathbb{E}(Y_i^0|D_i=1) - \mathbb{E}(Y_i^0|D_i=0)}_{ ext{Selection bias}} + \underbrace{(1-\pi)(ATT-ATU)}_{ ext{Heterogenous treatment effect bias}}$$

where $1-\pi$ denotes the share of people in the control group.

So there is a novel source of bias that comes from the potential *heterogeneity in the* individual treatment effect δ_i .

- *Selection bias*: those who attend university are likely to have higher baseline cognitive skills (regardless of whether they actually attend college).
- *Heterogeneous treatment effect bias*: those who attend university may improve their cognitive skills more at university because they are more motivated. back

