



# Variable Selection for Emulated Target Trials in Causal Survival Analysis

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# Motivation: Tamoxifen Duration in Early-Stage Breast Cancer

## Collaboration & Data Source

- Collaboration with **Gustave Roussy** and **Unicancer**.
- CANTO cohort (observational study with 9,000 patients and more than 4,000 variables).
- Study adjuvant **tamoxifen** therapy in breast cancer.

## Evidence from Clinical Trials

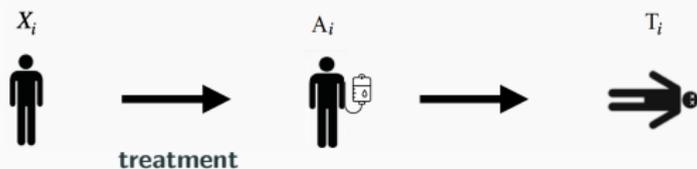
- Delozier et al. (2000): Longer tamoxifen treatment (>5 years) improved disease-free survival vs. stopping at 2–3 years.

## Real-World Perspective

- Trial inclusion criteria do not fully represent the real-world patient population.
- Clinicians are interested in assessing whether real-world evidence aligns with trial findings.

Aim: **treatment deescalation** without compromising patient outcomes.

# Causal Survival Analysis in observational study



$$\Rightarrow n \text{ i.i.d. } ( \underbrace{X_i}_{\text{covariates}}, \underbrace{A_i}_{\text{treatment}}, \underbrace{T_i}_{\text{outcome}} ) \in \mathbb{R}^d \times \{0, 1\} \times \mathbb{R} \times \mathbb{R}$$

Let's say that in our example  $X_1 = \text{sex}$ ,  $X_2 = \text{age}$  and  $A = 1$  (resp.  $A = 0$ ) denotes 5-year (resp. 3-year) treatment.

Covariates		Treatment	Censoring	Status	Outcomes		
$X_1$	$X_2$	<b>A</b>	<b>C</b>	$\Delta$	<b>T(0)</b>	<b>T(1)</b>	$\tilde{T}$
1	24	1	?	1	?	200	200
2	52	0	?	1	100	?	100
1	33	1	200	0	?	?	200

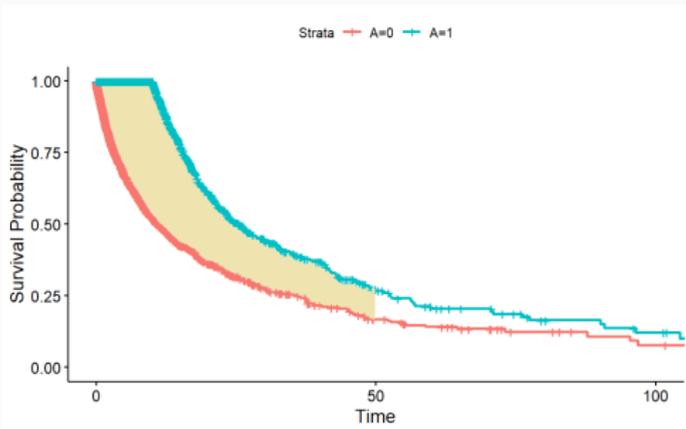
- $\Delta = \mathbb{1}\{C < T\}$  is the censoring indicator.
- $\tilde{T} = \min(T, C)$  is the observed outcome.

# Estimand : Difference of RMST

## Difference in RMST : Average treatment effect in survival analysis

$$\begin{aligned}\theta_{RMST}(\tau) &= E[\min(T(1), \tau) - \min(T(0), \tau)] \\ &= \int_0^{\tau} (S_1(t) - S_0(t)) dt\end{aligned}$$

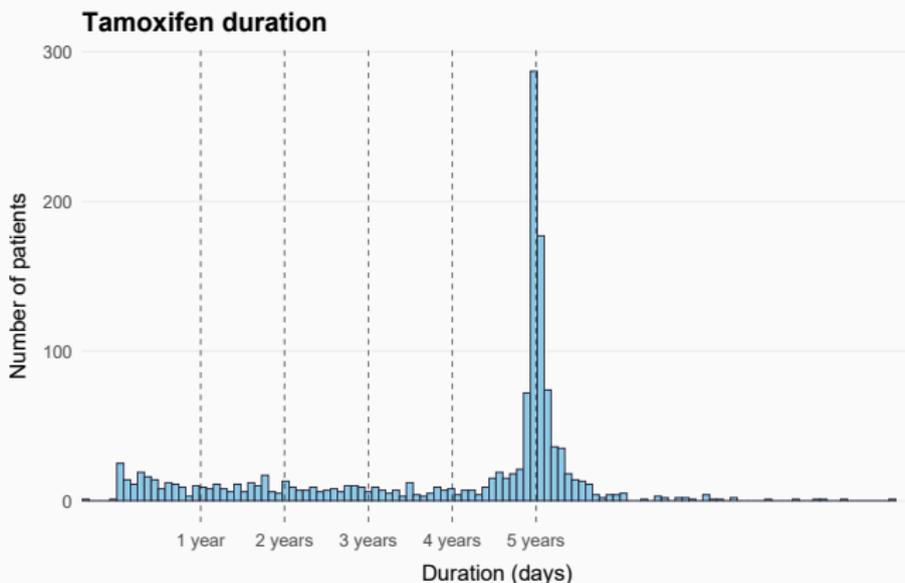
**RMST** can be defined as a measure of average survival from time 0 to time  $\tau$  a **fixed time horizon**



$\theta_{RMST}(\tau = 50) = 10$  means that on average the treatment increases the survival time by 10 days at 50 days.

# From Baseline Treatment Assignment to Treatment Duration

- Prior work assumed a baseline “3 vs 5 years” assignment.
- In our data, no such baseline assignment: treatment is **continuous**.
- Observed duration is **post-baseline**.



# From Baseline Treatment assignment to Treatment Duration

In CANTO settings, treatment is **not a one-time decision**:

## Post-baseline Classification

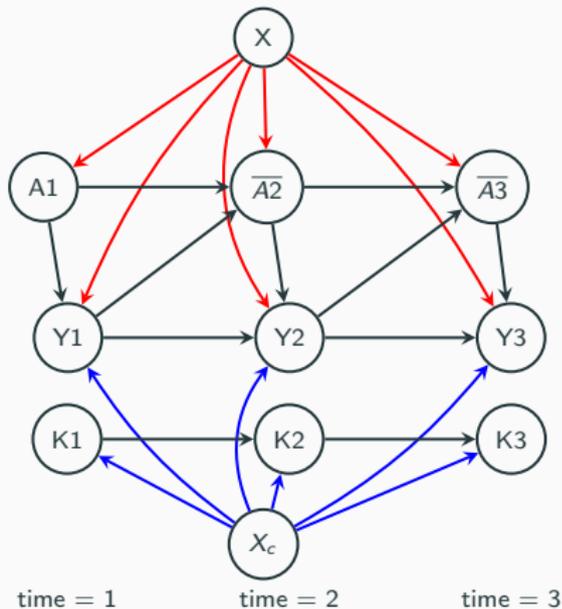
- Classification into “short” vs “long” duration is made **post-baseline**, based on follow-up data.

## Time-varying Dependence

Duration is **inherently linked** to survival:

- To have a long duration, the patient must survive long enough (**immortal time**).
- Survival also depends on treatment duration.

**Dynamic treatment assignment**  
(Real World assignment)

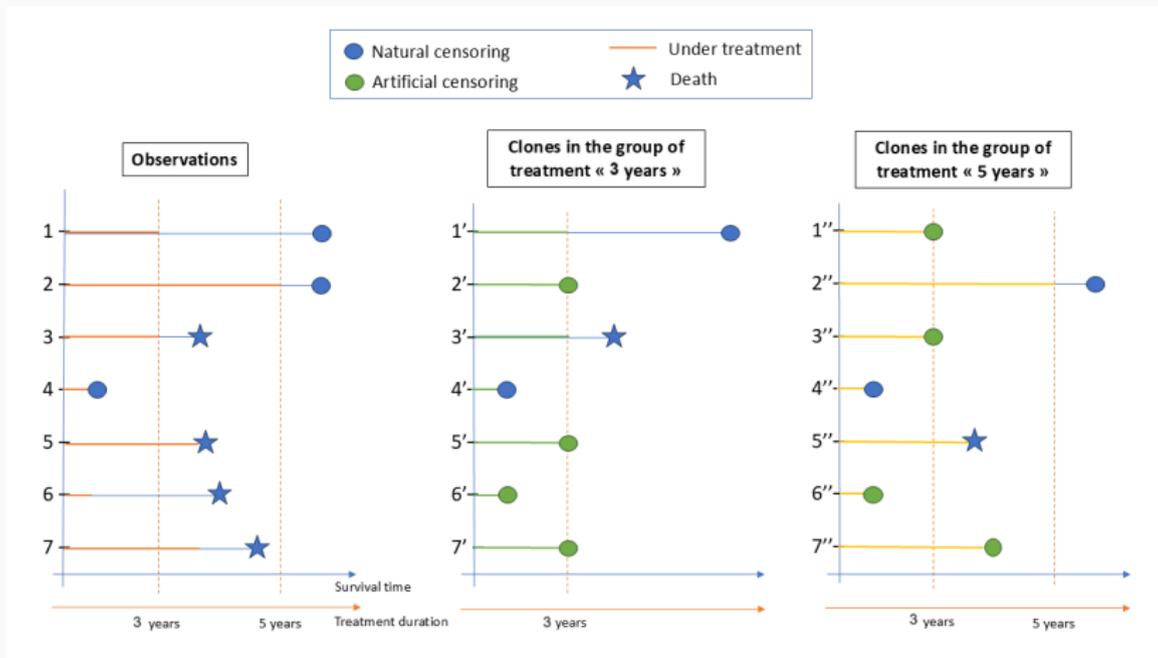


$K_t$  is the indicator of censoring,  $Y_t$  is an indicator of the event of interest,  $A_t$  the treatment variable,  $X$  the baseline confounders and  $X_c$  the variables of dependent censoring.

# Strategy for 3 years vs 5 years of treatment

## Cloning–Censoring–Weighting (Cain et al., 2010)

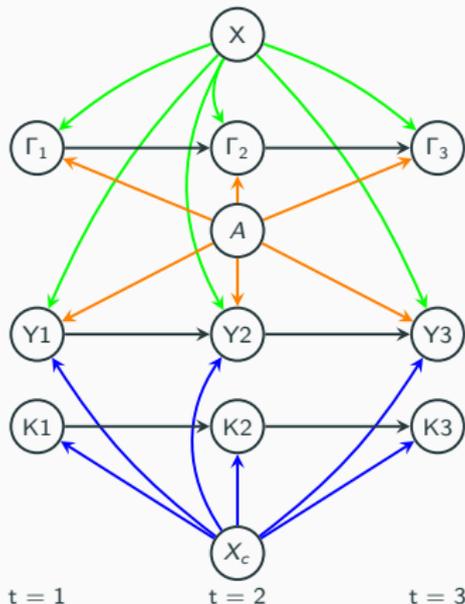
Clone each subject for each strategy; *artificially censor* at deviation:



# Strategy for 3 years vs 5 years of treatment

## Cloning–Censoring–Weighting (Cain et al., 2010)

$G$  is artificial **informative censoring** ( $G = \min(t \in [\tau], \Gamma_t = 1)$ ),  $C$  is natural **informative censoring** ( $C = \min(t \in [\tau], K_t = 1)$ ),  $T$  is the time to event outcome ( $T = \min(t \in [\tau], Y_t = 1)$ ).



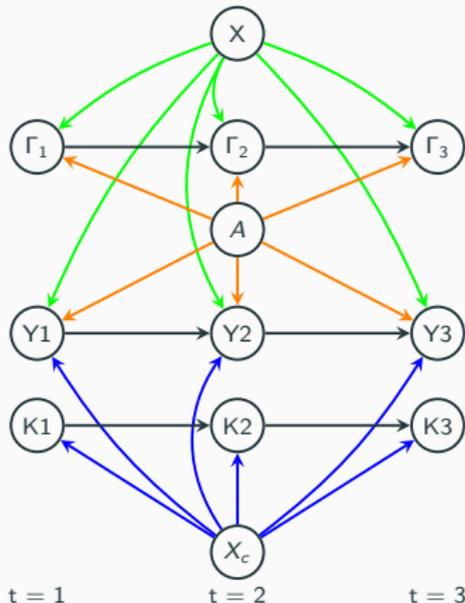
DAG after cloning ( $A=0$  if under first strategy and  $A=1$  if under second strategy).

$\Gamma_t$  is the indicator of artificial censoring.

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## Weighted Kaplan–Meier<sup>1</sup>:

Two informative censorings  $\Rightarrow$  compute weights

- **Artificial censoring:**

$$W_{i,a}^A(t) = \frac{\mathbf{1}\{G_i \geq t\}}{\prod_{j=0}^t \Pr[\Gamma_{i,j}=0 | \Gamma_{i,j-1}=0, Y_{i,j}=0, X, A_i=a]}$$

- **Natural censoring:**

$$W_{i,a}^N(t) = \frac{\mathbf{1}\{C_i \geq t\}}{\prod_{j=0}^t \Pr[K_{i,j}=0 | K_{i,j-1}=0, Y_{i,j}=0, X_c, A_i=a]}$$

Then weighted KM by  $w_{i,a}(t) = W_{i,a}^A(t) \cdot W_{i,a}^N(t)$  and compute  $\theta_{RMST}$ , the difference of RMST.

<sup>1</sup> Identifiability: exchangeability, positivity, consistency.

<sup>2</sup> Adjust on  $X_c$  and  $X$

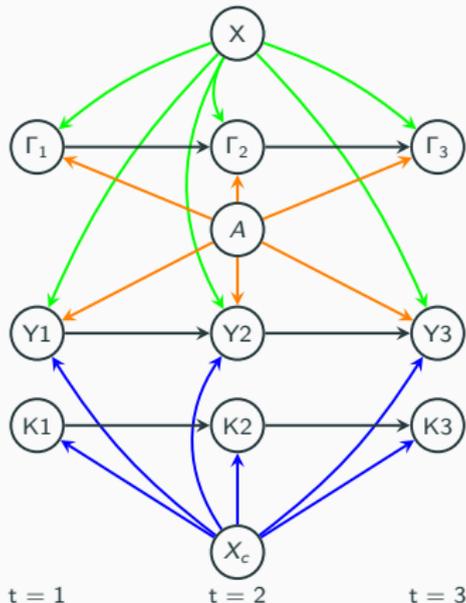
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### G-formula estimator<sup>1 2</sup>:

$$\hat{\theta}_{G\text{-formula}} = \frac{1}{n} \sum_{i=1}^n \hat{\mu}(x_i, 1) - \hat{\mu}(x_i, 0).$$

with  $\hat{\mu}(x_i, a) = \int_0^\tau S(t | X = x_i, A = a) dt$  the integral of the **conditional survival function** truncated at  $\tau$ .

<sup>1</sup> Identifiability: exchangeability, positivity, consistency.

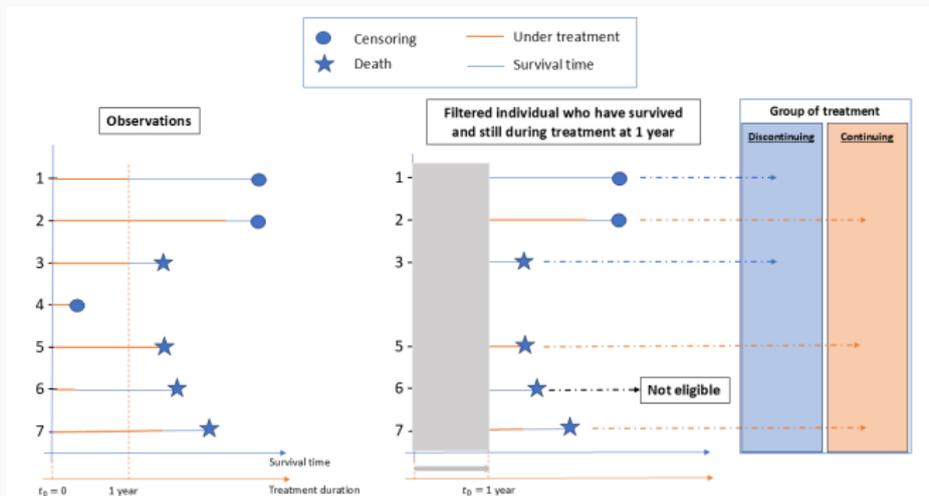
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# Strategy for finding optimal treatment duration

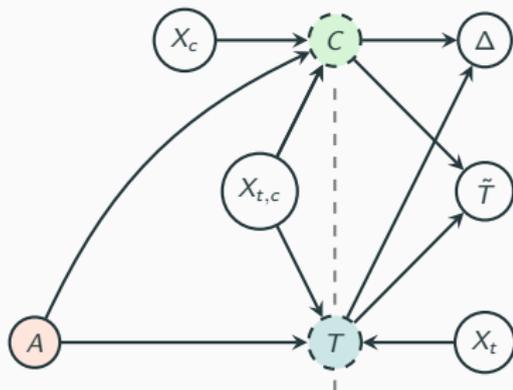
Evaluate at different sequential time  $t_1$ , among patients under treatment, compare stop vs continue treatment to learn the **optimal duration**.

## Target Trial Emulation

1. Specify the *target trial* (e.g. time-zero  $t_0 \in \{1, 2, 3, 4, 5\}$  years; eligible patients are those alive and on treatment at  $t_0$ , target 10-year survival)
2. Decide whether a **grace period** (may lead to immortal time bias) is used to classify individuals into “stop” vs “continue”.



# RCT with dependent censoring: Variable Selection



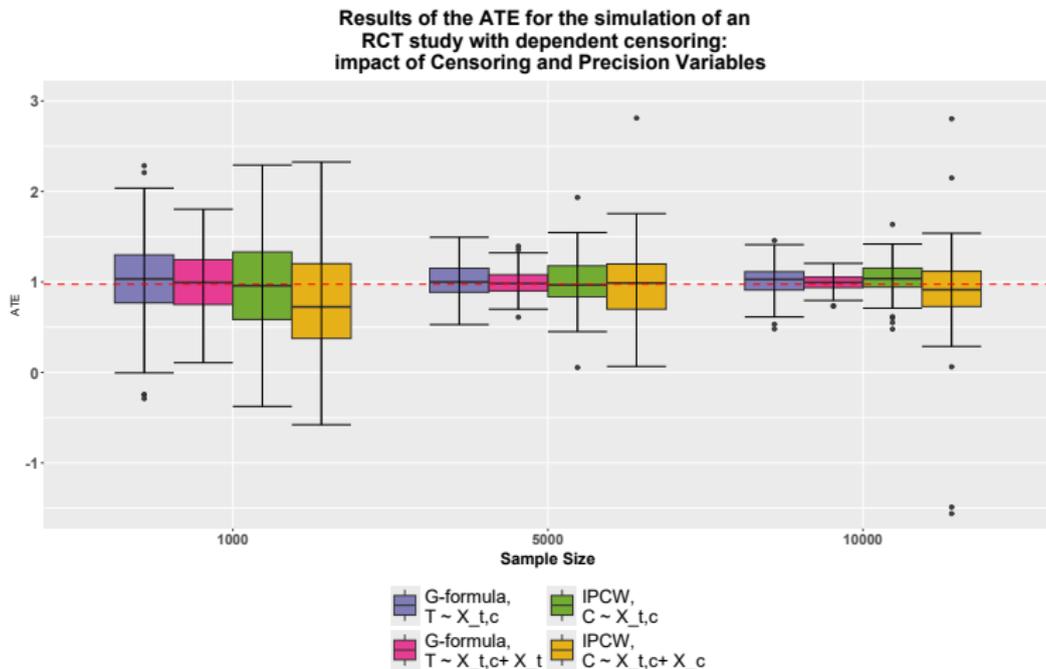
DAG of a RCT with dependent censoring ( $A$  is the treatment,  $T$  is the time to event outcome and  $C$  is the censoring time).

To ensure causal identifiability in RCT and dependent censoring, adjust each model for:

- **Censoring model and Outcome model:**  $X_{t,c}$  (remove selection bias from informative censoring)

**However:** Adding extra variables (precision, or censoring-related) can affect estimator variance.

# Baseline Treatment Assignment: Variable Selection



The simulations show that:

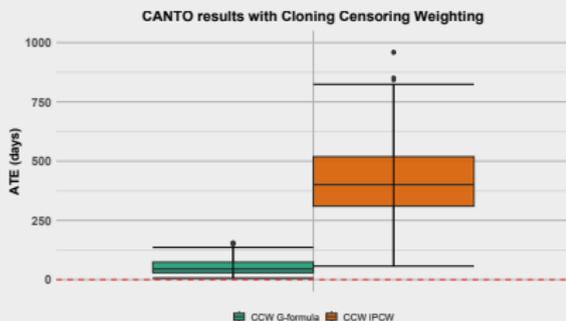
- Precision variables ( $X_t$ ) (Purple vs Pink): decrease the variance.
- Censoring-related variables ( $X_c$ ) (Green vs Yellow): increase the variance.

# RMST Difference Between Tamoxifen Durations

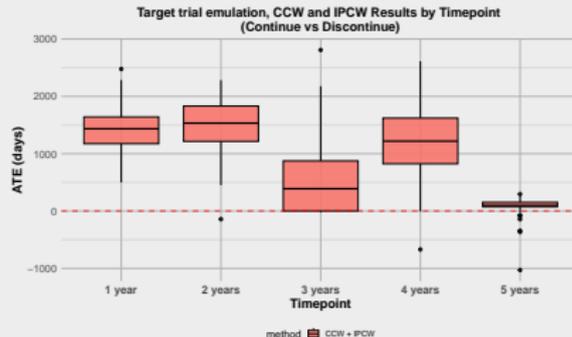
## Setup

- **Objective:** Estimate the ATE on 10-year OS of **3-year vs 5-year** treatment duration.
- **Grace period:** 6 months after (and/or before) the timepoint.
- **Adjustment:** baseline confounders only (no time-varying covariates).
- **Uncertainty:** boxplots over **150** bootstrap samples.

## 3 years vs 5 years



## Continue vs Discontinue



## Results

- **ATE** > 0 across analyses, better 10-year OS under the 5-year strategy.

# Conclusion & Perspectives

## Theoretical

- *Prove* the variance increase from including censoring-related variables in the censoring model.
- *Formalize* the Cloning–Censoring–Weighting approach.

## Data Analyses

- Incorporate *time-varying covariates* (longitudinal history).
- Identify potential *precision covariates*.

## Challenges

- Treatment dates often uncertain (missing day and/or month).
- Time-varying covariates are difficult to construct (continuous variables, choice of evaluation timepoints).

Thank you for your attention !  
Here's a summary of our previous contribution on practical  
recommendations for causal survival estimators:



Arxiv article (submitted)



Github repository



Feel free to give me some feedback or contact me for any question:  
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Bernard Sebastien and Charlotte Voinot are Sanofi employees and may hold shares and/or stock options in the company. Julie Josse have nothing to disclose.

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## References

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# Appendix



# Dynamic G-formula (no time-dependent confounding)

**Context.** We compare two *static treatment duration strategies* (e.g., 3y vs 5y). Since there is no time-dependent confounding, only **baseline covariates**  $X$  need to be adjusted for.

## Principle.

- Fit survival models  $P(T | A, X)$  within each interval of follow-up.
- Create counterfactual datasets where treatment trajectories are **forced**:

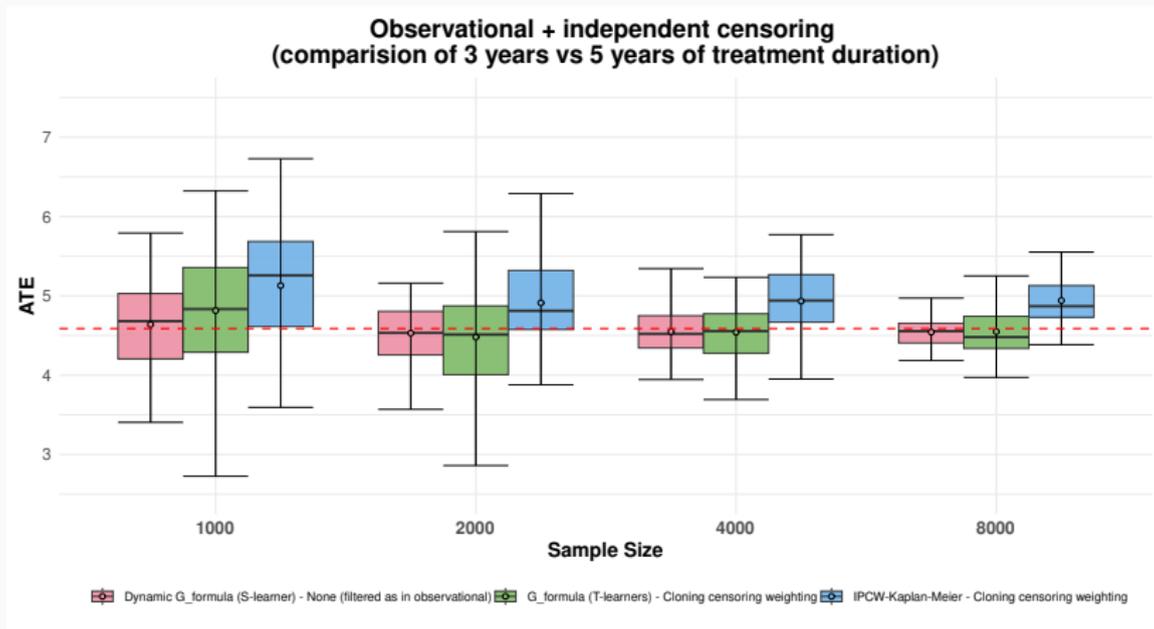
$$\bar{A}^{(0)} = (1, 1, 1, 0, 0), \quad \bar{A}^{(1)} = (1, 1, 1, 1, 1).$$

- Use fitted models to predict survival for all individuals under each strategy.
- Aggregate to obtain survival curves  $S^{(0)}(t), S^{(1)}(t)$  and functionals (e.g., RMST difference).

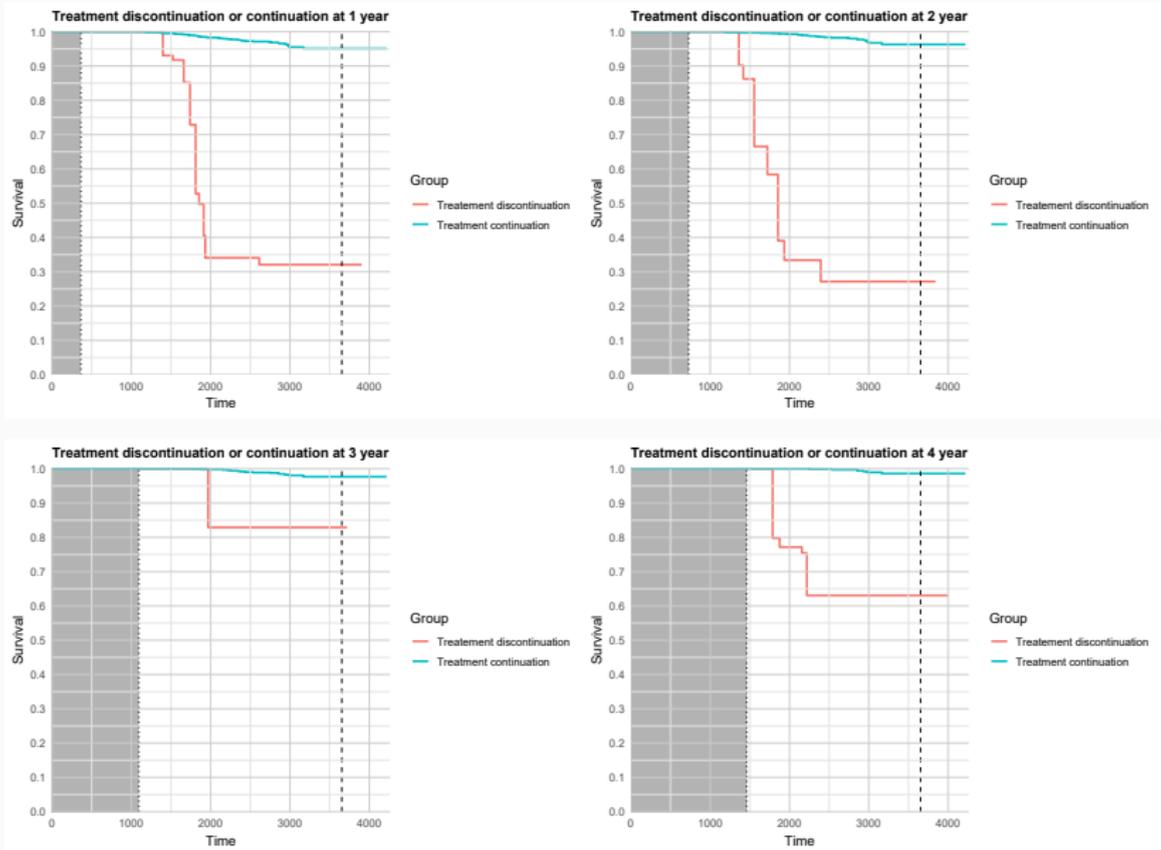
## Why dynamic g-formula here?

- Still framed longitudinally (assign  $\bar{A}_t$  at each visit),
- but no need to model covariate evolution since  $L_t$  does not depend on  $A$ .

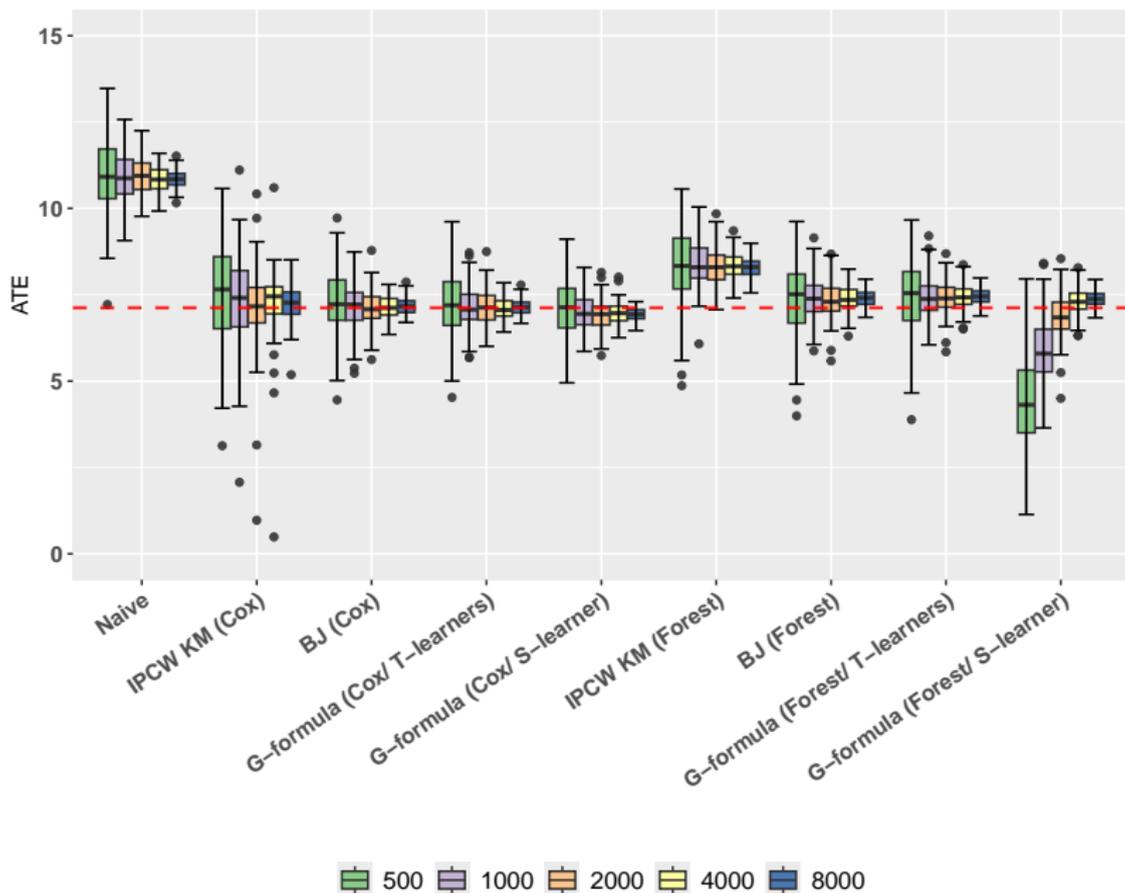
# Comparison of strategies to counter immortal time bias



# Survival curves of Tamoxifen duration: continue vs discontinue



# Estimators consistent in RCT with dependent censoring



# Why using RMST instead of Hazard Ratio ?

## Limitations of the Hazard Ratio (HR):

- Assumes **proportional hazards**.
- Hazard ratios **are not causally interpretable** (even in RCT unless there is no treatment effect) due to a built-in selection bias. [Martinussen and Vansteelandt (2013); Martinussen et al. (2020)]
- **Not collapsible**. [Huitfeldt et al. (2019); Greenland et al. (1999)]
- Difficult to interpret clinically. [Hernán (2010)]

## Advantages of Restricted Mean Survival Time (RMST):

- **Causal measure** as it is the extension of Average Treatment Effect. [Royston and Parmar (2013)]
- Does not necessarily require the assumption of proportional risks.
- Provides an **absolute measure**.
- **Clinically interpretable**: "How much longer does a patient survive on average within a given time frame?"

## Built-in selection bias HR

Here is the definition of Hazard ratio:

$$\exp(\beta) = \frac{\log P(T^1 > t)}{\log P(T^0 > t)}$$

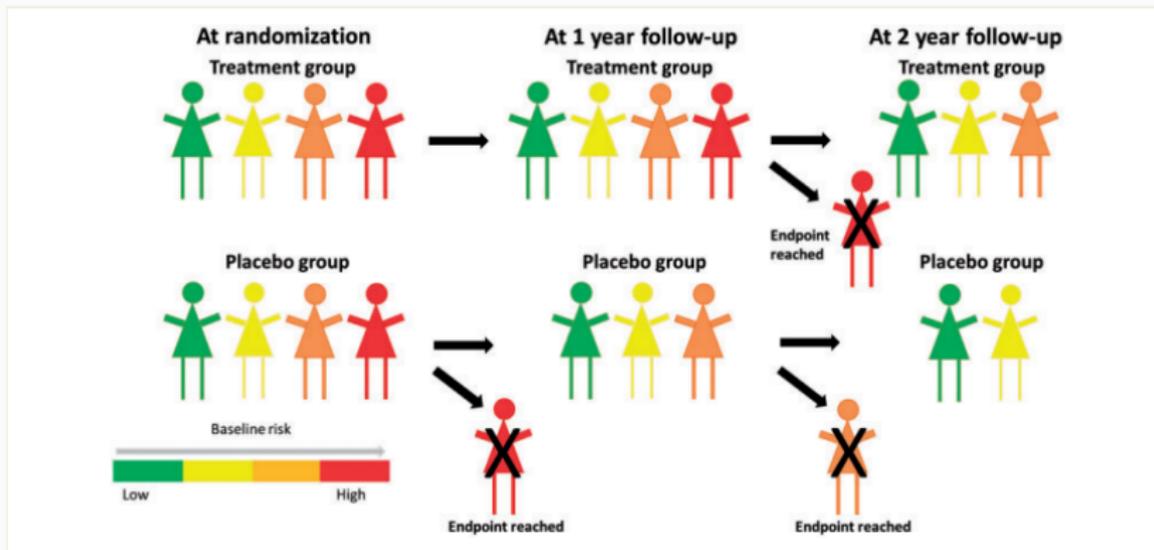
Under proportional hazard assumption, it goes:

$$\begin{aligned}\exp(\beta) &= \frac{\lim_{h \rightarrow 0} P(t \leq T < t+h | T \geq t, A=1)}{\lim_{h \rightarrow 0} P(t \leq T < t+h | T \geq t, A=0)} \\ &= \frac{\lim_{h \rightarrow 0} P(t \leq T^1 < t+h | T^1 \geq t)}{\lim_{h \rightarrow 0} P(t \leq T^0 < t+h | T^0 \geq t)}\end{aligned}$$

First, the right-hand expression shows that  $\exp(\beta)$  contrasts the hazard functions **with** and **without** intervention for two separate groups of individuals who survive time  $t > 0$ .

If it exists a treatment effect then, population **with** and **without** intervention cannot be compared anymore.

# Built-in selection bias HR



**Figure 2:** This schematic drawing illustrates the built-in selection bias in population-level hazard ratios: by definition, the population-level hazard ratio at a given time point is based on individuals who survived up to that time point, thereby it is a comparison between the unbalanced groups (from Stensrud et al. (2019))